

Spherical data validation of rock discontinuities orientation from Drone-derived 3D Point Clouds

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ABSTRACT

This work presents a methodology for the statistical validation of discontinuity surfaces obtained from point clouds using digital photogrammetry from drones. Our methodology allows you to review the quality of the data obtained with photogrammetry and decide whether these measurements are representative of the discontinuity surfaces that they analyze. It consists of three steps, the first one being a shape analysis that allows defining which statistical model should be used: Fisher for circularly symmetric clusters or Bingham fits better for axially symmetric clusters. This step also makes the most significant difference to other works since our methodology starts from the premise that not all discontinuity surfaces are flat. Therefore, Fisher parameters do not allow validating data that do not correspond to a plane.

In the second step of the methodology, we calculate the consistency parameters that depend on the statistical model defined in step 1. The parameters are similar for both models; both estimate κ which indicates how much the sample is concentrated around the mean orientation and validates the existence of this and which is the value of the generating angle of a cone with a 95 % confidence limit that it contains within the mean orientation.

Finally, step 3 is used when there are control measurements to compare the point cloud data and define if both samples characterize the same discontinuity surface in the rock mass.

The results obtained on a rock outcrop allowed us to observe that the measurements obtained from the drone faithfully represent the discontinuity surface analyzed when these were compared with the measurements made manually with the compass. Furthermore, the dispersion parameters (κ and α_{95}) yielded results that make it possible to ensure that 1) there is a preferential direction (mean orientation) and 2) the mean orientation is representative of the entire measured surface.

Key words: rock-discontinuities; spherical statistics models; drone; validation statistical; point cloud.

RESUMEN

Se presenta una metodología para la validación estadística para superficies de discontinuidad obtenidas de nubes de puntos utilizando fotogrametría digital a partir de drones. Esta metodología permite revisar

la calidad de datos obtenidos con fotogrametría y decidir si estas mediciones son o no representativas de las superficies de discontinuidad que analizan. Esta metodología se compone de tres pasos, siendo el primero un análisis de forma que permite definir qué modelo estadístico se debe utilizar: Fisher para agrupaciones circularmente simétricas o Bingham que se ajusta mejor para agrupaciones axialmente simétricas. Este paso también es el elemento que marca la mayor diferencia con respecto a otros trabajos ya que nuestra metodología parte de la premisa de que no todas las superficies de discontinuidad son planas y, por lo tanto, los parámetros de Fisher no permiten validar datos que no corresponden a un plano

En el segundo paso de la metodología, se calculan los parámetros de consistencia que dependen del modelo estadístico que se definió en el paso 1. Los parámetros son similares para los dos modelos, en ambos se calcula κ que indica qué tanto se concentra la muestra alrededor de la orientación media y valida la existencia de esta y α_{95} que es el valor del ángulo generador de un cono con un límite de confianza del 95 % de que contenga dentro la orientación media.

Por último, el paso 3 se utiliza cuando se tienen medidas de control con las que comparar los datos de la nube de puntos y permite definir si ambas muestras caracterizan a la misma superficie de discontinuidad en el macizo rocoso.

Los resultados obtenidos en una pared rocosa permitieron observar que las mediciones obtenidas a partir del dron representan fielmente a la superficie de discontinuidad analizada, cuando estas se compararon con las mediciones realizadas manualmente con la brújula. Además, los parámetros de dispersión (κ y α_{95}) arrojaron resultados que permiten asegurar que 1) existe una dirección preferencial (orientación promedio) y 2) que esta dirección es representativa de toda la superficie medida.

Palabras clave: discontinuidades; modelos de estadística esférica; dron; validación estadística; nubes de puntos.

INTRODUCTION

Orientation data of rock strata and their discontinuities constitute a key knowledge element to characterize rock masses, which is crucial for many geoscientific and engineering applications. In structural geology, for instance, statistical analysis of orientation data allows reconstructing the history and evolution of an acting stress field (Ramsay and Huber, 1987; Davis *et al.*, 2011; Allmendinger, *et al.* 2011). In exploration

geology, rock lineaments provide insights into different geometric patterns of ore deposits, oil/gas reservoirs, and fractured aquifers (Lerche, 1997; Fetter, 2001). In geotechnical engineering, dip and dip direction records are key inputs for developing slope-rock stability studies (Goodman, 1989; Hudson and Harrison, 1997).

To obtain rock strata and their discontinuities' orientation datasets, the common practice requires a qualified professional to visually identify and acquire an exhaustive number of measurements by hand, using a geological compass. Indeed, such process is time-consuming, labor-intensive and could be subjective because the measurement depends on person expertise. In recent years, local-scale remote sensing techniques, such as drone photogrammetry and Light Detection and Ranging (LiDAR), have emerged as promising tools for measuring rock discontinuities, which results particularly beneficial to acquire data over inaccessible or unsafe areas (Sturzenegger and Stead, 2009; Li, *et al.*, 2019; Buyer, *et al.*, 2020). Remote-sensing-derived orientation datasets are commonly processed programmatically with the assistance of segmentation algorithms, hence lacking in-field discrimination offered by an experienced professional. Nevertheless, such datasets are large enough to be assessed by the means of statistical approaches, thus providing robust outcomes with confidence intervals (Jordá Bordehore *et al.*, 2017; Drews *et al.*, 2018; Mancera-Alejándrez *et al.*, 2020).

In the field of Geosciences, orientation data considers the representation of lines and planes on a semi-sphere using a stereographic projection (Mardia, 1972; Fisher *et al.*, 1987; Priest, 1993; Marshak and Mitra, 1988; Davis, 2002; Borradaile, 2003). Overall, the analysis of spherical data using statistical methods relies mainly on assessing the consistency parameters given by several models. The Fisher Model (Fisher, 1953), for instance, assumes symmetrical data distribution around a mean orientation vector, whereas the Bingham Model (Bingham, 1964), focuses on distributions rather elliptical. The former is suitable for managing planar rock surfaces, while the latter is more convenient from treating non-planar discontinuities, *i.e.* curve surfaces (Woodcock, 1977).

The most used model for analyzing point clouds is Fisher's because it is generally convenient to simplify discontinuity surfaces to planes and most algorithms to extract surfaces from point clouds have been developed for analyzing planes (Riquelme *et al.* 2014; Borrmann *et al.* 2011; Rusu and Cousins 2011; Leng *et al.* 2016). The use of tools such as photogrammetry or LiDAR, allows it to be possible to obtain the real shape of the discontinuities, so it makes sense to explore those statistical distribution models that are coupled to it, such as Bingham's to curved surfaces.

For example, in the analysis for discontinuities' rock-slope stability, define whether it is a plane or curved surface, it has implications in calculating the factor of safety (FS) when it used limit equilibrium analysis. Since on the one hand, the models that are generally used are based on plane surfaces (Goodman, 1989; Hudson and Harrison, 1997), which is a simplification of reality (like plane and wedge sliding model's), and this sometimes could be translated into an underestimation or an overestimation of factor of security (FoS) (Sturzenegger and Stead, 2009; Stead and Wolter, 2015). On the other hand, although there is the complexity of mapping the real shape with measurements by hand, the use of techniques such as remote sensing generates numerous measurements of data that allow obtaining more realistic surfaces even if they are curved, so that in these cases the Fisher model will not always be applicable, and it will be necessary to have alternatives for the validation of the measurements obtained with remote sensing techniques.

Previous works perform statistical validations for point-cloud-derived orientation datasets using Fisher's model for symmetrical distributions. Jordá Bordehore *et al.* (2017) make the first approach

at methodically analyzing remote-sensing spherical data using one of the two dispersion parameters of Fisher's model: the concentration parameter κ for comparing compass *versus* photogrammetric data obtained using Structure from Motion (SfM) discontinuity surfaces. Later on, Drews *et al.* (2018) utilizes two of Fisher's dispersion parameters κ and α_{95} , the latter being a confidence parameter, for comparing orientation measurements from compass and LiDAR cloud points. Mancera *et al.* (2020) determines a one-order-of-magnitude improvement in the confidence cone parameter α_{95} , when using orientation data from a dense point cloud ($n = 2000$) *versus* limited compass measurements ($n = 36$) on a planar surface of rock mass. However, and to the best knowledge of the authors, no previous works heretofore have dealt with the spherical statistics analysis of point-cloud-derived orientation data lacking rotational symmetry around the mean orientation, rather fitting Bingham's model. Furthermore, no previous work has dealt with the characterization process of rock walls to determine whether Fisher or Bingham's Model are more suitable.

In this work, we propose a methodology for choosing the most convenient model between Fisher's and Bingham's and to perform a quality assessment of remotely-sensing-derived dip and dip direction measurements using spherical statistics and compass measurements. The proposed methodology can be divided into three steps: 1) Exploratory analysis of the point cloud data for determining the data distribution in their stereographic projection to define the most convenient distribution model (Fisher or Bingham). 2) Quality assessment based on the classical consistency parameters for spherical statistics (κ and α_{95}) depending on the model chosen in the first step, and 3) a hypothesis test for determining if the cloud point-derived data represents the same population as control measurements (in this case data obtained using a geologist compass). This methodology can serve as a standard for characterizing point clouds and providing a data quality measurement and give consistence parameters for review the dispersion of data.

The scientific contributions of the research are a simple and reproducible methodology that consists of three steps where it is possible to decide which statistical distribution model best suits the type of surface being measured and a quantitative assessment of the quality of data obtained with remote sensing such as digital photogrammetry or LiDAR.

THEORETICAL BASE FOR SPHERICAL DATA DISTRIBUTION MODELLING

In stereographic projections, the representation of a discontinuity plane is done using points known as poles. When a group of poles share similar distribution, they can be represented as a mean orientation pole. The reliability of a mean orientation to represent a discontinuity can be estimated using spherical statistics and distribution models, as well as consistency parameters. In Earth Sciences, two models are widely accepted (Davis, 2002; Borradaile, 2003).

The first model, proposed by Fisher (Fisher, 1953), focuses on the distribution of poles symmetrically distributed around a mean vector (Figure 1a). The model conceptualizes a concept called Fisher's spherical Probability Distribution Element (PDE), which provides the probability of finding a given direction within a unitary area element (dA), (Table 1). The model also considers a parameter known as Fisher's κ (*kappa*), which increases as the concentration of poles around the mean orientation increases (Figure 1b).

The second model, proposed by Bingham (Bingham, 1964), focuses on the populations of data forming elongated clusters in the form of girdles. The model defines a confidence cone about the mean orientation, which intersection with the semi-sphere of the stratigraphic projection

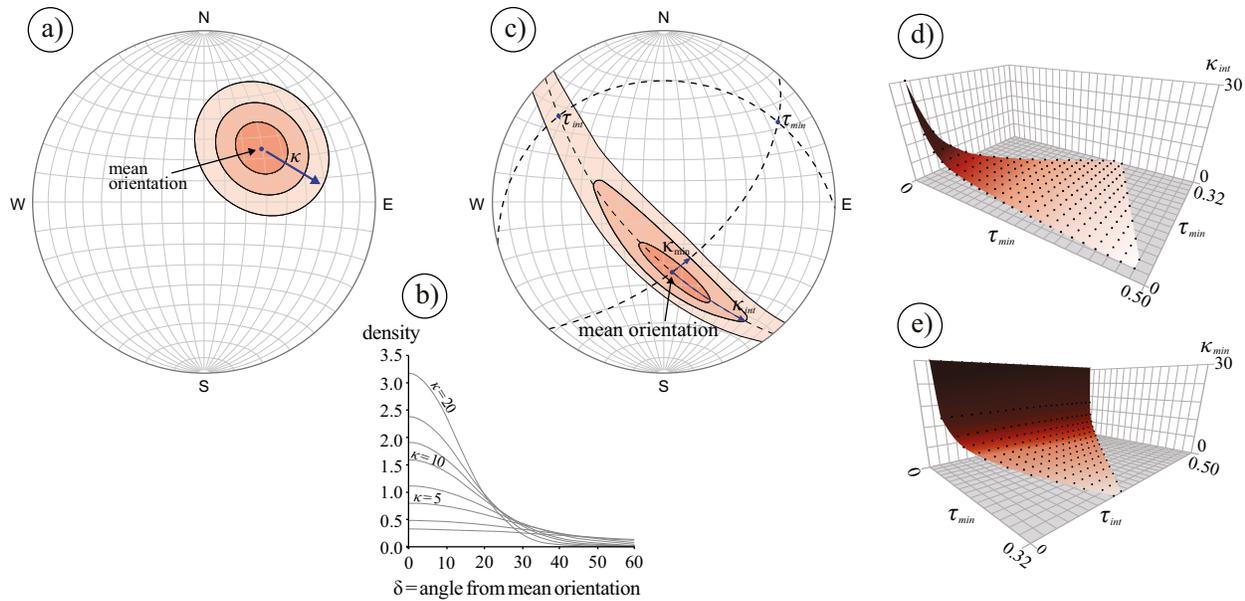


Figure 1. Distribution models: a) Fisher distribution: circular-symmetrical cluster of orientations. Its confidence cone is circular on the sphere. b) Probability density of Fisher distribution for different Fisher κ vs. the angular distance (δ) from mean orientation. c) Bingham distribution has a cluster of orientations that does not possess circular symmetry but rather an orthorhombic symmetry. Density contours range in shape from ellipses to complete great-circle girdles. d) and e) show the concentration parameters variation (κ_{int} and κ_{min}) in function of normalized eigenvalues τ_2 and τ_3 . Modified from (Borradaile, 2003).

tends to form an ellipse (Figure 1c). The distorted region controlled the concentration parameters κ_{int} and κ_{min} , they are intermediate and minimum concentration respectively and both are function of the normalized eigenvalues of the orientation's tensor τ_2 and τ_3 (Figure 1d and 1e). Bingham's spherical PDE is defined in Table 1.

METHODS

Methodology proposed

We propose a methodology to validate dip and dip direction data obtained from photogrammetry-derived point clouds. Such methodology is presented in Figure 2 and described in the following paragraphs. The methodology consists of three tests: the first two tests (“exploratory analysis” and “data consistency parameters estimation”) perform a quality check on the data, while the third test (“accuracy test”) compares the dip and dip direction data from photogrammetry against control measurements gathered using instruments such as a geological compass.

Exploratory analysis

This test aims at graphically identifying the distribution model that better fits the data and thence calculate the dispersion parameters. For this test, the pole distribution is plotted on an equal-area stereographic projection. The following step is realize a shape analysis using a Woodcock's diagram (Woodcock, 1977).

Woodcock (1977) proposes representing each observation in terms of its direction cosines l_i , m_i and n_i . The set of all data's directions can be represented as a 3 by 3 orientation matrix, T (Equation 1), by using x_i , y_i and z_i , which are individual observation's components without normalization in the X, Y and Z directions (Fernández, 2005) (1).

$$T = \begin{bmatrix} \sum l_i^2 & \sum l_i m_i & \sum l_i n_i \\ \sum l_i m_i & \sum m_i^2 & \sum m_i n_i \\ \sum l_i n_i & \sum m_i n_i & \sum n_i^2 \end{bmatrix} \tag{1}$$

Considering that T is a symmetrical matrix, its eigenvalues, $\lambda_1, \lambda_2, \lambda_3$, and its eigenvectors (v_1, v_2, v_3) can be calculated. The eigenvector

Table 1. Theoretical distributions of orientations of Fisher and Bingham models (Fisher, Lewis and Embleton, 1987; Borradaile, 2003).

Theoretical statistical model	Stereogram appearance	Probability density element PDE = $Ce^T dA$	T
Fisher (1953)	Circular-symmetric cluster around of mean orientation	$\frac{\kappa}{4\pi \sinh \kappa} e^T dA$	$\kappa[\sin\theta\sin\bar{\theta} \cos(\phi-\bar{\phi})] + \kappa\cos\theta\cos\bar{\theta}$
Bingham (1964) Kent (1982)	Elliptical clusters (girdle) about of mean orientation	$C_\kappa e^T dA$ where $C_\kappa \approx \frac{e^{-\kappa} (\kappa^2 - 4\beta^2)^{\frac{1}{2}}}{2\pi}$	$\kappa\chi'\xi_1 + \beta\chi'\xi_2 - \beta\chi'\xi_3$ where $\chi' = (\sin\theta\cos\phi, \sin\theta\sin\phi, \cos\theta)$

κ : concentration parameter; β : shape parameter = $\kappa_{int}/\kappa_{min}$; $\bar{\theta}, \bar{\phi}$ (polar coordinates): mean co-latitude, mean longitude, (ξ_1, ξ_2, ξ_3) is an orthogonal triple of unit vectors; the 3x3 orthogonal matrix can be regarded as a rotation matrix.

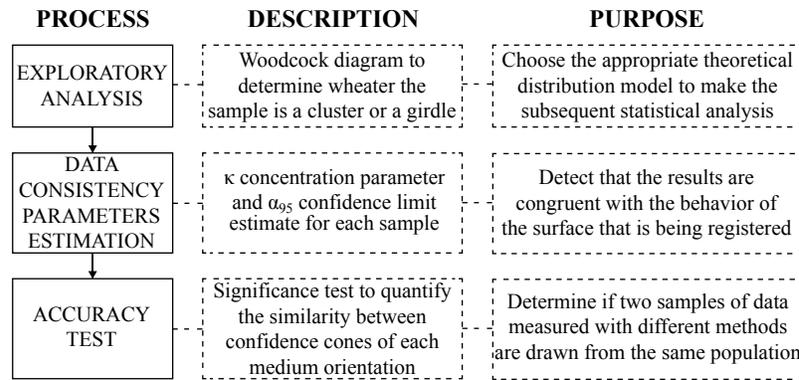


Figure 2. Flow chart for the implementation of the three tests conforming the methodology proposed in this work.

v_1 corresponds to the direction of the smallest inertial moment, and thus, corresponds to the orientation with maximum density of observations. Meanwhile, the eigenvector v_3 corresponds to the largest inertial moment and corresponds to orientation with minimum observations' density (Woodcock, 1977; Fernández, 2005). Expectedly, v_2 is perpendicular to eigenvectors v_1 and v_3 .

The sum of the eigenvalues λ_1 , λ_2 and λ_3 equal the number of observations of the analyzed set (N), Equation 2 (Woodcock, 1977):

$$\lambda_1 + \lambda_2 + \lambda_3 = N \tag{2}$$

after normalization, we obtain that:

$$\tau_1 + \tau_2 + \tau_3 = 1 \tag{3}$$

where: $\tau_1 = \lambda_1/N$

The normalized eigenvalues are ingested into a plot whose coordinate system has the comparison between τ_1 and τ_2 in the vertical axis and the comparison between τ_2 and τ_3 in the horizontal axis, forming Woodcock's diagram (Figure 3). On the one hand, the degree of fit to a plane is indicated by the parameter M (Equation 4). The larger the M , the more coplanar the poles, regardless of their symmetry around a point (girdle) or around an axis (cluster).

$$M = \ln \left(\frac{\tau_1}{\tau_2} \right) \tag{4}$$

On the other hand, the degree of collinearity is related with the parameter K (Equation 5). When K equals to 1, there's not a clear criterium to define a cluster or a girdle. When 1, the data set falls into the classification of a girdle, and when ∞ , the data set falls into the classification of a cluster.

$$K = \frac{\ln(\tau_1 / \tau_2)}{\ln(\tau_2 / \tau_3)} \tag{5}$$

Classic works on the treatment of paleomagnetic data (Tarling, 1983; McElhinny and McFadden, 2000; Tauxe, 2003; Butler, 2004) use the eigenvalues obtained from the direction matrix of the point cloud to be analyzed and plot them in Woodcoks diagram to obtain the model that better fits the data (cluster or girdle) (Woodcock, 1977). Thereafter, the best probabilistic approach is adopted, *i.e.* Fisher's for clusters and Bingham's for girdles. In this work, we propose adopting the same approach but for discontinuities' orientation data.

Data Consistency Parameters

Fisher's model parameters. Fisher (1953) defines κ and the confidence cone α_{95} . The first parameter, κ , is a concentration parameter

and describes the dispersion of the orientations around of the preferred orientation (mean orientation). κ served to calculate the PDE according to Table 1. A simple approximation of κ valid for any $\kappa \geq 3$ is given by Fisher's k :

$$\kappa \approx k = \frac{N-1}{N-R} \tag{6}$$

Where N is the sample size and R is the magnitude of the non-normalized resultant vector. R can be calculated by first obtaining the coordinates of each orientation datum in the X, Y, Z coordinate system by them dip and dip direction data:

$$x_i = \sin(dip) \cdot \sin(dipdir) \tag{7}$$

$$y_i = \sin(dip) \cdot \cos(dipdir) \tag{8}$$

$$z_i = \cos(dip) \tag{9}$$

Then, the sum of all elements over each axis can be calculated as:

$$x = \sum_{i=1}^N x_i; y = \sum_{i=1}^N y_i; z = \sum_{i=1}^N z_i \tag{10}$$

which are finally used to calculate R :

$$R = \sqrt{x^2 + y^2 + z^2} \tag{11}$$

The second parameter, α_{95} , is a 95 % confidence cone within which the true mean direction will fall, and can be calculated as (Tarling, 1983; McElhinny, McElhinny and McFadden, 2000; Borradaile, 2003):

$$\alpha_{95} = \frac{140^\circ}{\sqrt{kN}} \tag{12}$$

Both κ and α_{95} are accuracy measurements of the estimated mean orientation. According to the literature κ values greater than 10 are acceptable (Mardia, 1972; Tarling, 1983; Davis, 2002). As for α_{95} the smaller the value, the better.

Bingham model's parameters. As opposed to Fisher's model, Bingham's distribution model is not symmetric in all directions, but has a distortion in the confidence region around the mean. The parameters that define the confidence cone's shape, therefore, are κ_{int} and κ_{min} , which are perpendicular to the largest concentration one, κ_{max} . κ_{int} and κ_{min} relate to eigenvalues τ_2 and τ_3 according to the following expressions (Mardia and Zemroch, 1977):

$$\int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} \left[e^{\left\{ \frac{1}{2}(\kappa_{min} + \kappa_{int}) \sin^2 \theta \right\}} \right] \left[\pi (\sin^3 \theta - 2\tau_3 \sin \theta) I_0 \left\{ \frac{1}{2}(\kappa_{min} - \kappa_{int}) \sin^2 \theta \right\} + \sin^3 \theta I_1 \left\{ \frac{1}{2}(\kappa_{min} - \kappa_{int}) \sin^2 \theta \right\} \right] d\theta = 0 \tag{13}$$

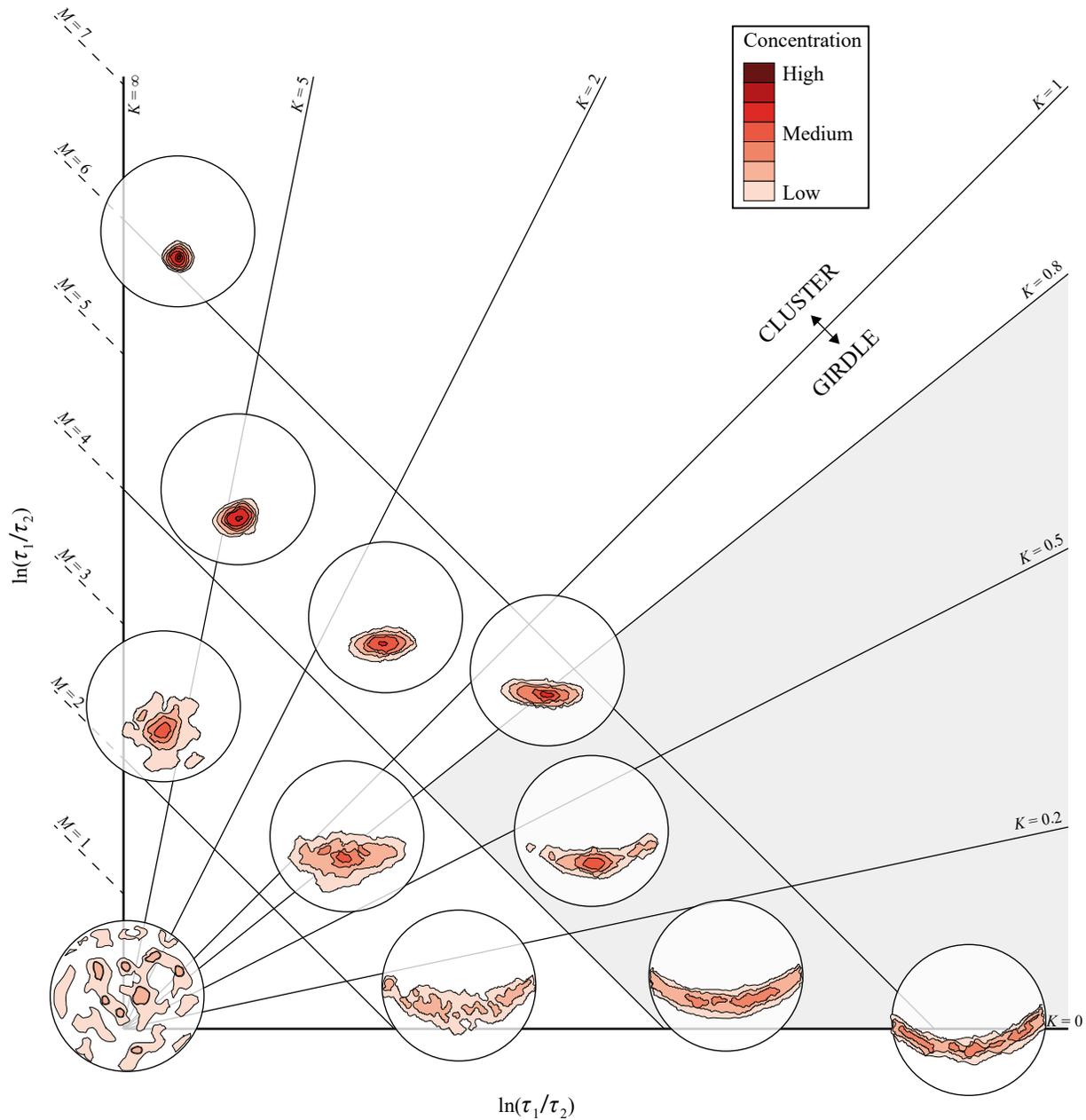


Figure 3. Woodcock's diagram showing the cluster's (circularly symmetric) and gridle's (axially symmetric) fields. (Modified from Woodcock, 1997). Observe that the definition of Cluster or Girdle depends on the coordinate of the data points with respect to line $K=1$. Where $K = \ln(\tau_1/\tau_2)/\ln(\tau_2/\tau_3)$, and $M = \ln(\tau_1/\tau_3)$.

$$\int_2^{\frac{\pi}{2}} \left[e^{\frac{1}{2}(\kappa_{min} + \kappa_{int}) \sin^2 \theta} \right] \left[\pi (\sin^3 \theta - 2\tau_2 \sin \theta) I_0 \left\{ \frac{1}{2}(\kappa_{min} - \kappa_{int}) \sin^2 \theta \right\} + \sin^3 \theta I_1 \left\{ \frac{1}{2}(\kappa_{min} - \kappa_{int}) \sin^2 \theta \right\} \right] d\theta = 0 \quad (14)$$

Where I_0 and I_1 are modified Bessel's functions of the first kind and orders zero and one respectively, τ_2 and τ_3 are derived from Equations 2 and 3, and θ is the cone angle. Mardia y Zemroch (1977) provide a table with the solutions to Equations 13 and 14 (Table 2) for values of $\kappa_{min} < 0.05$. If necessary, for small values of τ_2 and τ_3 , Equation 15 provides an acceptable approximation.

$$\kappa_i = \frac{1}{2\tau_i}; \quad i = int, min \quad (15)$$

Finally, from the concentration values κ_{min} and κ_{int} we can obtain the precision of the orientations in the minimum and intermediate direction (Borradaile, 2003) (Equations 16 and 17).

$$\alpha_{95}(int) = \frac{140^\circ}{\sqrt{\kappa_{int} N}} \quad (16)$$

$$\alpha_{95}(min) = \frac{140^\circ}{\sqrt{\kappa_{min} N}} \quad (17)$$

Accuracy test

The accuracy test is applicable when control measurements are

Table 2. Concentration parameters, κ_{min} and κ_{int} for the Bingham (Mardia and Zemroch, 1977).

$\tau_3 \backslash \tau_2$	0.02	0.04	0.06	0.08	0.10	0.12	0.14	0.16	0.18	0.20	0.22	0.24	0.26	0.28	0.30	0.32
0.02	25.55															
0.04	25.56	13.11														
0.06	25.58	13.14	9.04													
0.08	25.60	13.16	9.07	7.04												
0.10	25.62	13.18	9.08	7.04	5.80											
0.12	25.63	13.19	9.09	7.04	5.79	4.92										
0.14	25.64	13.20	9.09	7.03	5.77	4.90	4.23									
0.16	25.65	13.20	9.08	7.02	5.75	4.87	4.20	3.66								
0.18	25.65	13.19	9.07	7.00	5.73	4.84	4.16	3.62	3.16							
0.20	25.64	13.18	9.05	6.97	5.69	4.80	4.13	3.57	3.11	2.71						
0.22	25.63	13.17	9.03	6.94	5.66	4.76	4.07	3.52	3.05	2.65	2.29					
0.24	25.61	13.14	9.00	6.91	5.62	4.71	4.02	3.46	2.99	2.58	2.22	1.89				
0.26	25.59	13.12	8.97	6.87	5.57	4.66	3.97	3.40	2.93	2.52	2.15	1.81	1.50			
0.28	25.57	13.09	8.93	6.83	5.52	4.61	3.91	3.34	2.86	2.44	2.07	1.72	1.41	1.11		
0.30	25.54	13.05	8.89	6.78	5.47	4.55	3.84	3.27	2.79	2.36	1.98	1.63	1.31	1.00	0.71	
0.32	25.50	13.01	8.84	6.73	5.41	4.48	3.77	3.20	2.71	2.28	1.89	1.54	1.21	0.89	0.59	0.29
0.34	25.46	12.96	8.79	6.67	5.35	4.42	3.70	3.12	2.62	2.19	1.79	1.43	1.09	0.77	0.46	0.15
0.36	25.42	12.91	8.73	6.61	5.28	4.34	3.62	3.03	2.53	2.09	1.69	1.32	0.97	0.64		
0.38	25.37	12.86	8.67	6.54	5.21	4.26	3.35	2.94	2.43	1.99	1.58	1.20				
0.40	25.31	12.80	8.60	6.47	5.13	4.18	3.45	2.85	2.33	1.87						
0.42	25.25	12.73	8.53	6.39	5.05	4.09	3.35	2.74								
0.44	25.19	12.66	8.45	6.31	4.96	3.99										
0.46	25.12	12.58	8.37	6.22												
0.48	25.04	12.50														

κ_{int} 25.61
 κ_{min} 2.43

available from a different, more “truer” source, to compare against the measurements from the point clouds throughout the evaluation of the degree of confidence cones overlap. An informal but quick analysis consists of observing if one confidence cone falls within the other; if both confidence cones do not overlap at all, the hypothesis considering that both samples belong to the same population is immediately rejected for a significance level of 0.05 (McElhinny and McFadden, 2000). When two samples’s confidence cones overlap but do not contain

the mean direction, the *F*-test (*F*-statistic) is applied. *F* is defined as a function of both samples size (N_1 and N_2), the resultants for each one of the samples R_1 and R_2 , and the combined resultant *R* (Equation 11) (Watson, 1965; McElhinny and McFadden, 2000).

$$F = (N_1 + N_2 - 2) \left[\frac{R_1 + R_2 - R}{N_1 + N_2 - R_1 - R_2} \right] \quad (18)$$

If the calculated *F*-value of Equation 18 exceeds to Tabled value

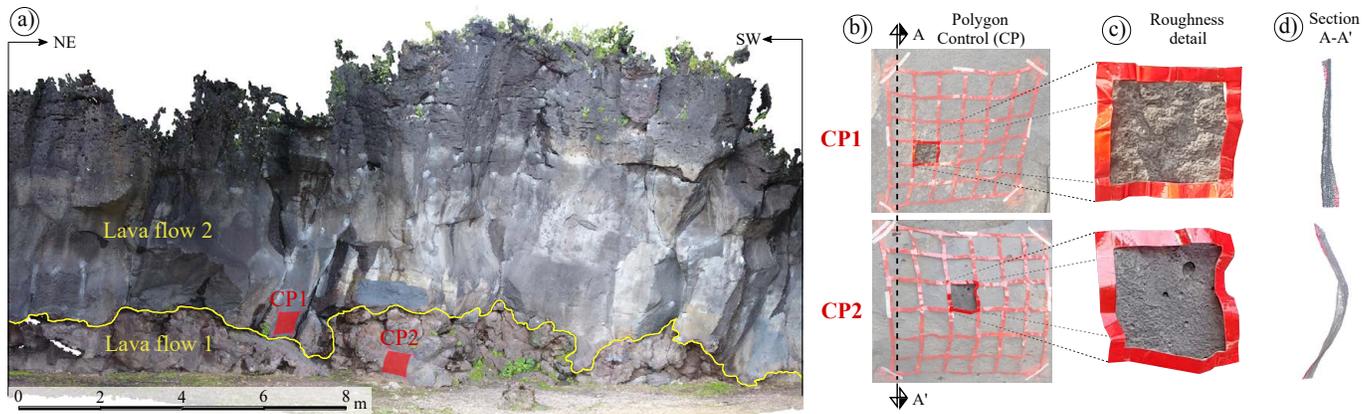


Figure 4. “La Escuelita” rock outcrop wall (19.3341°, -99.1933°). (a) Panoramic view, yellow polyline indicates a discontinuity surface between two lava flows and red polygons show the location Control Polygons 1 and 2. (b) Red meshes show the discretization, of CP1 corresponds to a planar discontinuity surface, while CP2 corresponds to a curved one. (c) Roughness detail for one element on the mesh. (d) Show a section of planar surface (CP1) and curved surface (CP2).

(Supplementary Table S1), the null hypothesis that the samples of mean orientations are similar or represent the same discontinuity surface is rejecting.

Experiment design and data acquisition strategy

Study area

We choose a rock wall, which is a basaltic rocks (Badilla, 1977) outcrop known as “La Escuelita” of 25.0 by 10.0 meters, located northwards from the Ciudad Universitaria at Olympic Stadium within UNAM’s main campus, Mexico City (19.3341°, -99.1933°) (Figure 4). It consists of a sequence of lava flows from the Xitle monogenetic volcano (Schmitter, 1953; Delgado *et al.*, 1998; Siebe, 2008). We use this outcrop because: a) the site is easily accessible for drone and manual measurements, b) the incipient vegetation in the wall to allow high-quality drone photogrammetry results, c) the outcrop presents several distinguishable discontinuities, from well-defined planes to oblique surfaces, mostly originated from lava-flow cooling.

Manually-acquired orientation data

We acquired orientation data manually over two areas using a geologist compass. We chose two Control Polygons in the wall, namely CP1 and CP2, which correspond to a planar discontinuity surface (CP1) and an oblique discontinuity surface (CP2) (see location in Figure 4a). CP1 and CP2 were chosen to be 0.6 by 0.6 m and a mesh was over imposed on CP1 and CP2 to discretize them into 0.1 by 0.1 m cells (Figure 4b). In each one of the cells, dip and dip direction were measured using a geologist compass, retrieving 36 observations in total from each control polygon. The corners of each CP were later marked with red spray paint to make the area visible in the aerial pictures.

Drone-acquired orientation data

We generated point-wise orientation data from photogrammetry from an 3D-oriented pictures set acquired using a drone-carried gyrosopic camera. We used a DJI Phantom 4 Pro quadcopter with a 1” CMOS sensor of 20 megapixels and 100-12800 ISO range. We performed three wall-perpendicular flights with a distance of 8.0 m from the wall, 80% between-picture overlap, and 70 % between-line overlap, acquiring a total of 33 pictures (~ 160 MB) covering a 250 m² area. The pictures set was later ingested into AgisoftMetashape v. 1.5 (Agisoft, 2019) to conduct photogrammetry using 10 total-station-assisted ground control points placed on the wall’s surface, followed by an automated plane-detection process, assigning dip and dip direction to each point-

cloud element. Simultaneously, we computed each point’s direction cosines.

The photogrammetry output was of more than seven million points (Figure 4a). Later on, a subset of 2000 points within each control polygon were selected and cropped using Cloud Compare (Cloud Compare, 2019) (inset in Figure 4d) to compare against the data acquired manually.

RESULTS

Exploratory analysis of orientation data

We obtained the normalized eigenvalues for each control polygon (CP1 and CP2), which we summarize in Table 3.

We used the results of CP1 and CP2 (Table 3) to plot their coordinate in Woodcock’s diagram (Figure 5). We observe that for CP1 (triangle symbols) both point-cloud-derived and compass-acquired data sets yield results that fall close to the $K = 5$ line, well within the classification of a cluster. Therefore, CP1 datasets are analyzed using Fisher’s model. On the other hand, we observe that CP2 (square symbols) data sets yield results in the band $0.5 < K < 0.8$, in the girdle classification. Consequently, CP2 datasets are analyzed using Bingham’s model.

Data consistency parameters estimation and accuracy test

Control Polygon 1 (CP1)

In Section 4.3 we found that CP1 data defines a cluster and should be analyzed using Fisher’s model. Thus, we calculated κ y α_{95} consistency parameters, using equations 6 and 12, for both the manually-acquired data and the point-cloud subset. We present a summary of the parameters used for such calculations along with the accuracy test

Table 3. Eigenvalues and logarithm of eigenvalues ratios for CP1 and CP2.

Eigenvalues	CP1		CP2	
	Compass	Point Cloud	Compass	Point cloud
τ_1	0.9848	0.9839	0.8079	0.8316
τ_2	0.0114	0.0102	0.1793	0.1557
τ_3	0.0037	0.0059	0.0128	0.0127
$\ln(\tau_1/\tau_2)$	4.459	4.569	1.5054	1.6754
$\ln(\tau_2/\tau_3)$	1.125	0.547	2.6396	2.5063

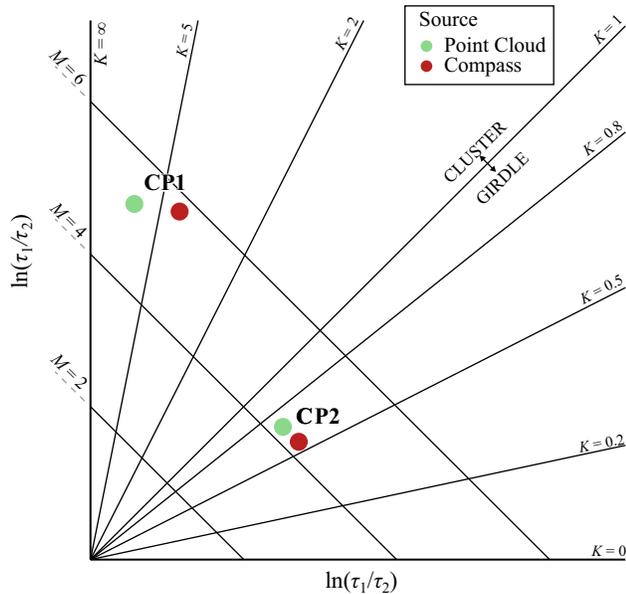


Figure 5. Control polygon 1 (CP1) and control polygon (CP2) on Woodcock's diagram.

results in Table 4. Fisher's κ , in both cases, is one order of magnitude larger than 10, which indicates a good concentration around the reported mean orientations (Mardia, 1972; Tarling, 1983; Davis, 2002). Such deduction can also be made from Figure 5, where we observe that both mean orientation points fall within the region defined by $M = 5$ and $M = 6$, and very close to $K = 5$. As for α_{95} , we obtain values of $< 1^\circ$ and 2° for the point cloud and the manual measurements, respectively. Thus, the mean orientation does accurately represent the presented discontinuity surface.

We present the data sets from both the manual measurements and the drone cloud, as well as the resulting mean orientation poles and confidence cones obtained from Table 4 data. Figure 6 shows that the point cloud's confidence cone is located inside the manual measure-

Table 4. Results summary of the dispersion parameters and accuracy test estimations for CP1 with Fisher's model.

CP1	Manual measurements	Point cloud
Mean orientation (Dip, Dip direction)	73.76°, 324.08°	74.03°, 325.31°
N	36	2036
R	35.72	2019.48
κ	127.39	123.17
α_{95}	2.07°	0.28°
<i>Accuracy test</i>		
Angular deviation	1.21	
F-test (95 %)	0.97 < 3	

ment's one. Moreover, we use the results from the F-test (Table 4) to ensure that both samples belong to the same discontinuity surface. We use the size from the manual measurements ($N_1 = 36$) and point cloud ($N_2 = 2036$) data sets in Equation 18 and compare it to the results from using two and 4140 degrees of freedom in Fisher's table (Supplementary Table S1). Because the F-test value calculated from our data, 0.97, is less than $2.9957 \sim 3$ obtained from Fisher's table, we conclude that both data sets indeed belong to the same surface.

Control Polygon 2 (CP2)

We found that CP2 data defines a Girdle and should be analyzed using Bingham's model (Figure 5). Hence, we calculated the dispersion parameters κ_{int} and κ_{min} and concentration parameters $\alpha_{95}(int)$ and $\alpha_{95}(min)$ in the minimum and intermediate eigenvalues' directions τ_2 and τ_3 to obtain the regions of confidence's semi-axes. We calculated κ_{int} and κ_{min} using τ_2 and τ_3 following Mardia and Zemroch (1977) approach. Furthermore, we used the simplified estimation for small τ_2 and τ_3 provided in Eq. 16. We present a summary of such results in Table 5, a stereographic representation of the data, mean direction and confidence cones in Figure 7a, and a zoom-in to the mean directions in Figure 7b.

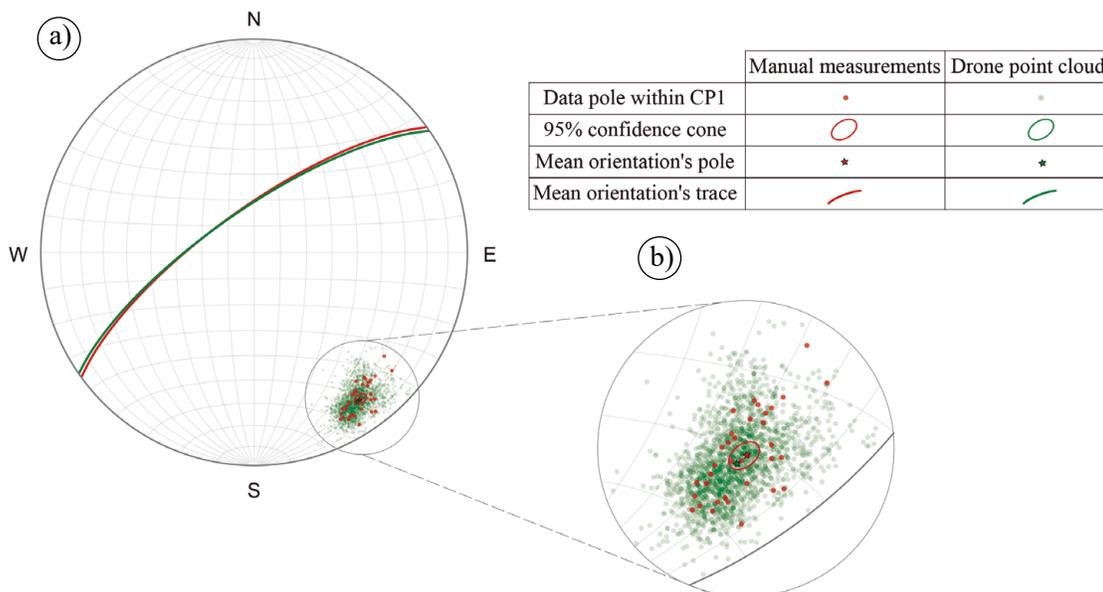


Figure 6. (a) Stereographical representation of CP1 data and (b) the confidence cones detail obtained using Fisher's model.

Table 5. Results summary of the dispersion parameters and accuracy test estimations for CP2 data using Bingham's model.

CP2	Manual measurements	Point cloud
Mean orientation (Dip, Dip direction)	83.89°, 5.08°	87.68°, 2.89°
N	35	2000
κ_{min}	43.62 ¹ /39.06 ²	43.87 ¹ /39.37 ²
κ_{int}	3.04 ¹ /2.79 ²	4.24 ¹ /3.21 ²
$\alpha_{95}(min)$	3.58°	0.47°
$\alpha_{95}(int)$	13.60°	1.52°
<i>Accuracy test</i>		
<i>F-test</i> (95 %)	0.004 < 0.103	

We observed a low dispersion from the κ_{min} value, which is, once again, greater than the threshold value of 10 (Tarling, 1983; Davis, 2002; McElhinny and McFadden, 2000) (*i.e.* $\kappa_{min} > 10$), and a considerably low $\alpha_{95}(min)$ value, especially for the point cloud data. The results of the accuracy test show that the confidence cones from both data sets overlap (7a y 7b). The *F-test* result for a 95 % confidence interval (0.004, see Table 5) reveals a value lower than the value obtained from the Fisher's table (0.013). Therefore, the hypothesis that both samples belong to the same sample is valid.

We calculate the dispersion parameters and perform the *F-test* on the CP2 data using Fisher's model, and show the results in Table 6, Figure 7c. Using Fisher's model, the calculated confidence cones for the manual measurements and the point cloud do overlap. However, the α_{95} values calculated for the manual measurements and the point cloud (32° and 5°) define a region evidently too wide (see Figure 7c) and far off the main concentration of poles. Furthermore, the results of the *F-test* are greater than the *F-value* from tables by two orders of magnitude, which would indicate that the samples do not belong to the same surface.

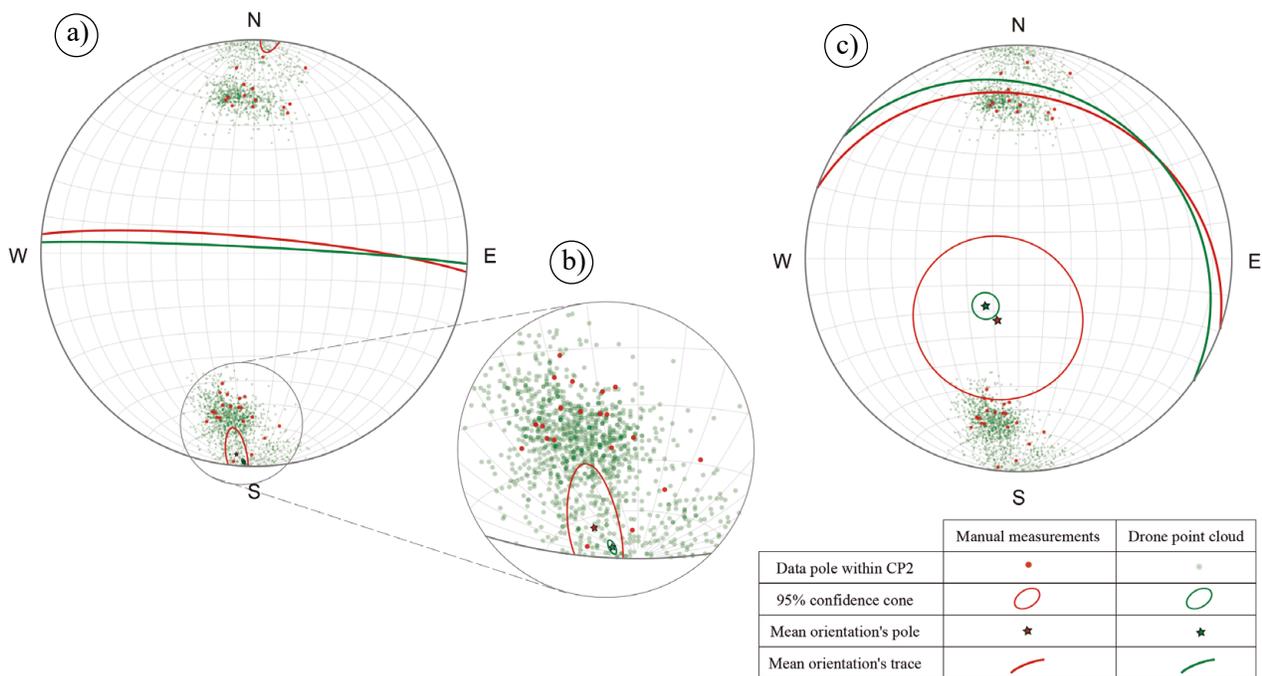


Figure 7. Stereographical representation of CP2 data and the calculated confidence cones. a) Results from using Bingham's model and b) a close up to the mean direction pole. c) Results from using Fisher's model.

DISCUSSION

The analysis of CP1 from Fisher's model revealed high κ values (Table 4), both for the compass data ($\kappa = 127.4$) and for the point cloud data ($\kappa = 123.2$), which indicates that the mean direction obtained for each case is representative of that sample. Studies from Jordá Bordehore *et al.* (2017) and Drews *et al.* (2018) showed cases where the condition of κ is also greater than 10, from 54 to 166 for the first and between 61 and 220 for the second (Table 7), which is consistent with our results. These studies also show that κ values can vary remarkably between orientation data obtained with a compass and point clouds for the same discontinuity surface (Table 7), indicating greater or lesser concentration but identifying a mean orientation, which is not the case of our results.

However, α_{95} values in our study differ by almost 2° between the manual and the point cloud measurements α_{95} (manual) = 2.07° and α_{95} (point cloud) = 0.28°, which means that the value of the mean orientation from the point cloud is of greater confidence. This contrast is observed when comparing the confidence cones' size in Figure 6b and is partly due to the large amount of data from the point cloud to capture the flat surface CP1.

For CP2, we used Bingham model. The values of κ_{min} and κ_{int} are almost identical for manual measurement, and the values of $\alpha_{95}(min)$ is 3.58 and $\alpha_{95}(int)$ is 3.6. These values define opened confidence cones compared to the results from the point cloud measurements, where are $\alpha_{95}(min)$ point cloud is 0.47 and $\alpha_{95}(int)$ point cloud is 1.52.

In contrast, point cloud measurements describe $\kappa_{min} > \kappa_{int}$ (Table 5), as expected according to Mardia and Zemroch (1977), since the sample clusters more in the minimum direction (see Figure 7b). As $\kappa_{int} < 10$, the data is scattered in that direction, but considering $\kappa_{min} > 1$, we can say that the sample has a preferential orientation. The value of α_{95} of greatest interest is $\alpha_{95}(min)$ since the confidence cone will be smaller in this direction.

As can be seen in Table 5 and Figure 7a, $\alpha_{95}(min)$ manual = 13.6° and $\alpha_{95}(min)$ point cloud = 1.52°, which indicates, as in the previous

CONCLUSIONS

In this work, a methodology is proposed to validate the quality of directional data such as point clouds or data acquired manually using a geological compass (contingent on sufficient data). The proposed methodology consists of three steps which are based on spherical statistics.

The first step is an exploratory analysis of shape to define which statistics distribution model to use. If the distribution of the data is circularly symmetric, the model to be used will be Fisher's and when the symmetry is axial in the data, then it is proposed to use the Bingham model. The second step is to calculate the consistency parameters of the data, that is, the concentration parameter (κ) and confidence cone (α_{95}). These parameters indicate how dispersed the sample is and the reliability of finding the mean orientation with a confidence limit of 95 %. The third step is a hypothesis test to determine if two groups of data represent the same surface in the mass rock.

We verify that the methodology allows obtaining parameters that validate the quality of the directional data and is easy to apply. Although the common professional practice is to treat discontinuities as planar surfaces to apply Fisher's model, we observed that there are surfaces that may resemble planes but statistically behave like non-planar surfaces which can be better represented using Bingham's model. The tendency to non-planarity is often considered as roughness and accounted for using parameters such as the Joint Roughness Coefficient (JRC) (Barton, 1973). However, such parameter works at a scale which usually does not impact the geometry used in a kinematic analysis of slope stability. Therefore, the choice of using either Fisher or Bingham's model becomes very important.

The limitations of the drone-assisted dataset used in this work includes the illumination of the observed discontinuities, in terms of the need to have a free face and must be free of vegetation. However, our dataset did allow us to map nonplanar discontinuities, which would be difficult to identify in a normal field survey without doing systematic measurements for 0.60 by 0.60 m polygons as done in this work.

SUPPLEMENTARY MATERIAL

The Table S1 "Distribution F-Fisher values" can be found at <www.rmccg.unam.mx>, in the abstract-preview page of this paper.

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